Variance Estimation Using Quality Characteristic

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Abstract This paper deals with estimation of unknown population variance of study variable (y) using auxiliary qualitative characteristic. The properties of the proposed class of estimators is studied and is supported through two real data sets.

Keywords: Mean square error; proportion; population variance; percent relative efficiency; simple random sampling without replacement.

1. INTRODUCTION

It is common in sampling to use auxiliary information to construct improved estimators. Ratio, product and difference estimators are common examples. One way to increase efficiency is the use of auxiliary quality characteristic. The common instance is-

- (a) Gender and length of the persons,
- (b) Yield of milk and variety of cow (see [3]).

Using point bi-serial correlation between the variables, many authors have suggested improved estimators for estimating unknown population mean (see [8] and [10]). Many a times in place of estimating population mean we are interested in estimation of unknown population variance (see [9]). Using known values of population parameter(s) Das and Tripathi [1], Isaki [2], Prasad and Singh [7], Kadilar and Cingi [4] and Singh et al. [10] have suggested estimators for estimating S_{y}^{2} .

Consider a finite population of size N with units $U = \{U_1, U_2, ..., U_N\}$.

Let y_i and ψ_i , be the values for y and ψ respectively (i=1,2,...,N). Let,

 $\psi_i = 1$, if ith unit of population has characteristic ψ and $\psi_i = 0$, otherwise. Let

$$M = \sum_{i=1}^{N} \psi_i$$
, $m = \sum_{i=1}^{n} \psi_i$, $P = \frac{M}{N}$ and $p = \frac{m}{n}$.

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Mishra, P. Singh, R.

We use following values to obtain value of mean squared error (MSE) $g_0 = \frac{s_y^2}{S_y^2} - 1$ and $g_1 = \frac{p}{P} - 1$, such that $E(g_i) = 0$; i = (0,1) and $E(g_0^2) = \xi(\lambda_{40} - 1), E(g_i) = 0, E(g_0g_1) = \xi C_p \lambda_{21}$. Where $\xi = \frac{1}{n} - \frac{1}{N}, \qquad C_p^2 = \frac{S_y^2}{P^2}, \qquad \lambda_{rs} = \frac{\mu_{rs}}{\mu_{20}^{r/2} \mu_{02}^{s/2}},$ $\mu_{rs} = \frac{1}{N-1} \sum_{i=1}^{N} (Y_i - \overline{Y})^r (\psi_i - P)^s$.

Consider estimator $H_0(=s_y^2)$ for estimation of S_y^2 -

$$H_0 = s_y^2 \tag{1}$$

The variance of H_0 is given by

$$Var(H_0) = \xi S_y^4 (\lambda_{40} - 1)$$
⁽²⁾

2. PROPOSED ESTIMATOR

A ratio estimator H_1 for estimating S_y^2 is given by

$$H_1 = s_y^2 \frac{P}{p} \tag{3}$$

Expression of mean squared error for H_1 is given by

$$MSE(H_1) = \xi S_y^4 \left\{ (\lambda_{40} - 1) + C_p^2 - 2C_p \lambda_{21} \right\}$$
(4)

Product Estimator: The conventional product estimator for estimating S_y^2 is given by

$$H_2 = s_y^2 \frac{p}{P} \tag{5}$$

Expression of mean squared error for H₂ is given by

$$MSE(H_2) = \xi S_y^4 \left\{ (\lambda_{40} - 1) + C_p^2 + 2C_p \lambda_{21} \right\}$$
(6)

Difference-Type Estimator: The usual unbiased difference-type estimator is given by

$$H_3 = s_y^2 + k(P - p).$$
(7)

where k is an unknown constant.

The variance of the estimator H_3 is given by

$$Var.(H_3) = \xi S_y^4 (\lambda_{40} - 1) + k^2 P^2 \xi C_p^2 - 2k P S_y^2 \xi C_p \lambda_{21}.$$
 (8)

At k $\left(=\frac{S_{y}^{2}\lambda_{21}}{PC_{p}}\right)$, the minimum value of the equation (8) is given by

$$Var_{\min}(H_3) = \xi S_y^4 (\lambda_{40} - 1 - \lambda_{21}^2).$$

3. ADAPTED ESTIMATOR

Motivated by Sahai and Ray [12], Koyuncu [5] and using the conventional product estimator of s_y^2 , we propose the variance estimator using auxiliary attribute p, as-

$$H_4 = \left\{ k_1 s_y^2 \left(2 - \frac{p}{P} \right)^{\gamma} \left(\frac{p}{P} \right)^{1-\gamma} + k_2 \left(\frac{p}{P} \right)^{\gamma} \right\} \exp\left\{ \frac{\eta \left(P - p \right)}{\eta \left(P + p \right) + 2\lambda} \right\}.$$
(10)

 k_1 and k_2 are constants. Here γ, η and λ are known population parameters of auxiliary attribute, like coefficient of variation C_p , population proportion P etc. (see [14]).

Expanding Eq. (10), we have

$$H_{4} = k_{1}S_{y}^{2} \begin{cases} 1 + g_{0} + (1 - \alpha - 2\gamma)g_{1} + \left(2\gamma^{2} + \frac{3}{2}\alpha^{2} - \alpha - 2\gamma + 2\alpha\gamma\right) \\ \times g_{1}^{2} + (1 - \alpha - 2\gamma)g_{0}g_{1} \end{cases} \\ + k_{2} \left\{ 1 + (\gamma - \alpha)g_{1} + \left(\frac{3}{2}\alpha^{2} - \gamma\alpha + \frac{\gamma^{2}}{2} - \frac{\gamma}{2}\right)g_{1}^{2} \right\}$$
(11)

here $\alpha = \frac{\eta P}{2(\eta P + \lambda)}$.

Solving (11), we get

$$MSE(H_4) = S_y^4 + k_1^2 A_1 + k_2^2 A_2 - 2k_1 A_3 - 2k_2 A_4 + 2k_1 k_2 A_5$$
(12)

Variance Estimation Using Quality Characteristic

(9)

Mishra, P. Singh, R. where,

$$\begin{split} A_{1} &= S_{y}^{4} \bigg[1 + \xi \bigg\{ \begin{pmatrix} \lambda_{40} - 1 \end{pmatrix} + C_{p}^{2} \left(1 + 4\alpha^{2} + 8\gamma^{2} - 4\alpha - 8\gamma + 8\alpha\gamma \right) \\ &+ 4C_{p}\lambda_{21} \left(1 - \alpha - 2\gamma \right) \bigg\} \bigg], \\ A_{2} &= \bigg[1 + \xi C_{p}^{2} \left(4\alpha^{2} + 2\gamma^{2} - \gamma - 4\alpha\gamma \right) \bigg], \\ A_{3} &= S_{y}^{4} \bigg[1 + \xi \bigg\{ C_{p}^{2} \bigg\{ \frac{3}{2}\alpha^{2} + 2\gamma^{2} - \alpha - 2\gamma + 2\alpha\gamma \bigg\} + C_{p}\lambda_{21} \left(1 - \alpha - 2\gamma \right) \bigg\} \bigg], \\ A_{4} &= S_{y}^{2} \bigg[1 + \xi C_{p}^{2} \bigg\{ \frac{3}{2}\alpha^{2} + \frac{1}{2}\gamma^{2} - \frac{1}{2}\gamma - \alpha\gamma \bigg\} \bigg], \\ A_{5} &= S_{y}^{2} \bigg[1 + \xi \bigg\{ C_{p}^{2} \bigg\{ 4\alpha^{2} + \frac{1}{2}\gamma^{2} - 2\alpha - \frac{3}{2}\gamma + 2\alpha\gamma \bigg\} + C_{p}\lambda_{21} \left(1 - 2\alpha - \gamma \right) \bigg\} \bigg] \end{split}$$

On partially differentiating Eq. (12) with respect to the constants, we get

$$k_1^{opt.} = \frac{A_2 A_3 - A_4 A_5}{A_1 A_2 - A_5^2}$$
 and $k_2^{opt.} = \frac{A_1 A_4 - A_3 A_5}{A_1 A_2 - A_5^2}$

Using optimum values of k_1 and k_2 in Eq. (12), we obtain

$$MSE_{\min}(H_4) = S_y^4 + \frac{2A_3A_4A_5 - A_2A_3^2 - A_1A_4^2}{A_1A_2 - A_5^2}$$
(13)

4. EFFICIENCY COMPARISONS

Comparison of estimators considered in this article-

From (2) and (9), $Var(H_0) - Var_{\min}(H_3) \ge 0$, if $\xi S_y^4 \lambda_{21}^2 \ge 0$. From (2) and (13), $Var(H_0) - MSE_{\min}(H_4) \ge 0$, if

$$S_{y}^{4}\left\{\xi\left(\lambda_{40}-1\right)-1\right\}-\frac{2A_{3}A_{4}A_{5}-A_{2}A_{3}^{2}-A_{1}A_{4}^{2}}{A_{1}A_{2}-A_{5}^{2}}\geq0.$$

From (9) and (13), $Var_{\min}\left(H_3\right) - MSE_{\min}\left(H_4\right) \ge 0$, if

$$S_{y}^{4}\left\{\xi\left(\lambda_{40}-1-\lambda_{21}^{2}\right)-1\right\}-\frac{2A_{3}A_{4}A_{5}-A_{2}A_{3}^{2}-A_{1}A_{4}^{2}}{A_{1}A_{2}-A_{5}^{2}}\geq0\,.$$

5. EMPIRICAL STUDY

In this section theoretical results are verified with help of following two data sets:

Variance Estimation Using Quality Characteristic

Population Data Sets:

| Population Data 1 | Population Data 2 |
|--|---|
| [Source: [6]] | [Source: [11]] |
| Y: Household size and | Y: Number of villages in the circles and |
| Ψ : Household took an agricultural loan from a bank. | $\Psi {:}~A$ circle consisting more than five villages. |
| N=25, n=8, P=0.4, $C_p^2 = 1.5625$, | N=89, n=23, P=0.1235, C_p^2 =7.7172, |
| $\lambda_{40} = 2.3077, \ \lambda_{21} = -0.16801, \ S_y^2 = 17.4233.$ | $\lambda_{40} = 3.8542, \ \lambda_{21} = 1.3228, \ S_y^2 = 4.0738.$ |

6. CONCLUSION

We have proposed some new estimators for estimating population variance using auxiliary quality characteristics. Tables 1, 2 and 3 reveal that the estimators proposed performs better than other estimators considered in the present study.

| Estimators | PRE (population 1) | PRE (population 2) |
|------------------|----------------------|----------------------|
| H ₀ | 100.00 | 100.00 |
| \mathbf{H}_{1} | 39.7453 | 88.5871 |
| H_2 | 53.3712 | 15.9268 |
| H ₃ | 102.2059 | 258.4389 |

| η | Λ | PRE ($\gamma = 0$) | PRE ($\gamma = 1$) |
|-------|-------|-----------------------------|-----------------------------|
| 1 | 1 | 7896.457 | 767.5542 |
| 1 | 0 | 727.0434 | 2388.9827 |
| C_p | 1 | 5819.49 | 811.4687 |
| 1 | C_p | 10944.1369 | 730.8589 |
| | | | |

| Mishra, P. | Р | 1 | 33706.1568 | 653.1480 |
|------------|-------|-------|------------|-----------|
| Singh, R. | 1 | Р | 2629.017 | 1001.8551 |
| | Р | C_p | 49789.2678 | 636.9053 |
| | C_p | Р | 2145.2016 | 1081.5429 |

| Table 3: Percent relative efficiency | of the j | proposed est | timator H ₄ 1 | for Population 2. |
|--------------------------------------|----------|--------------|--------------------------|-------------------|
|--------------------------------------|----------|--------------|--------------------------|-------------------|

| η | Λ | PRE ($\gamma = 0$) | PRE ($\gamma = 1$) |
|-------|-------|-----------------------------|-----------------------------|
| 1 | 1 | 369192.9438 | 1005.0804 |
| C_p | 1 | 87858.40 | 1172.6728 |
| 1 | C_p | 2247258.882 | 932.6347 |
| Р | 1 | 17438366.19 | 904.0739 |
| 1 | Р | 49156.0955 | 1482.9417 |
| Р | C_p | 130508750.5 | 894.3079 |

CONFLICT OF INTEREST

The author(s) declare(s) that there is no conflict of interest regarding the publication of this paper.

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Variance Estimation Using Quality Characteristic